

## TOBIT ANALYSIS USING SAS

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"What do you mean less than nothing?" replied Wilbur. "I don't think there is any such thing as less than nothing. Nothing is absolutely the limit of nothingness. It's the lowest you can go. It's the end of the line. How can something be less than nothing? If there were something that was less than nothing than nothing would not be nothing, it would be something--even though it's just a very little bit of something. But if nothing is nothing, then nothing has nothing that is less than it is."

E. B. White, *Charlotte's Web*  
(New York: Harper, 1952) p. 28

### INTRODUCTION

Tobit analysis, first proposed by James Tobin (5), estimates the parameters of a regression model when the dependent variable is limited by some upper or lower bound. This limit is often zero, as in the case of the amount annually expended by a family for an expensive durable good. Every year many families refrain from purchasing this good. Thus the limiting value of the dependent variable not only exists as a possibility, but is often observed. Under such conditions, Tobit estimation can be viewed as an extension of Probit analysis; families are not merely classified as purchasers or non-purchasers, but the magnitude of the expenditure is maintained as relevant.

Tobit estimates were originally computed by various modifications of Newton's method. These are discussed in Tobin (5) and Anemiyah (1). Ray Fair (3) proposed an alternative procedure which has reduced the computing time for many problems although it generally requires more iterations than the procedures based on Newton's method. Program steps and notation presented in this paper closely follow the procedure outlined by Fair.

### THE MODEL

The model is

$$y_t = \beta_0' X_t + u_t \text{ if RHS} > 0 \\ = 0 \text{ otherwise} \quad t = 1, 2, \dots, T$$

where  $\beta_0$  is a  $K \times 1$  vector of unknown coefficients,  $X_t$  is a  $K \times 1$  vector of values of the explanatory variables for observation  $t$ , and the error term  $u_t$  is independently distributed  $N(0, \sigma^2)$ .

$Y_t$  is positive for  $R$  observations and zero for  $S=T-R$  observations.

### THE PROGRAM

The SAS program presented in this paper iteratively calculates the Tobit estimates of  $\beta_0$  and  $\sigma^2$  using the MATRIX procedure in SAS. The estimation procedure is considered to have achieved convergence when the magnitude of the greatest change in the estimate of  $\beta_0$  on successive iterations is less than the tolerance level of .001 (SAS Statement 98). There is no guarantee that the procedure will converge and the program ceases processing if the convergence criterion has failed to be met in 100 iterations.

The authors are currently investigating the optimal starting values for  $\beta_0$ . This program begins with either a vector of zeros or the ordinary least squares estimates for  $\beta_0$  for the  $R$  observations with positive  $y_t$ . The vector of zeros is selected if the number of observations with  $y_t$  equal to zero exceeds the number of observations with nonzero  $y_t$  (SAS Statements 78 and 79).

Fair (3) suggests some value between .3 and .5 for the damping factor ( $\lambda$ ). This program initially sets the value of  $\lambda$  at .5 and reduces it at the rate of .05 for each successive round of 20 iterations (SAS Statement 83).

The SAS program presented in this paper outputs the number of iterations required for convergence, the variance-covariance matrix of  $(\beta_0, \sigma^2)$ , and the standard table of regression estimates with their standard errors. Sample output is displayed in Figure 1.

The matrix procedure outputs a SAS data set containing the observed  $Y_t$  values, the predicted  $y_t$  values, the residuals and the  $X_t$  values. This facilitates further investigation of the residuals via SAS graphic and analytical procedures and also allows comparisons with the residuals obtained using ordinary least squares methods.

#### ACKNOWLEDGEMENTS

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#### REFERENCES

1. Amemiya, Takeshi: "Regression Analysis When the Dependent Variable is Truncated Normal," Econometrica, 41(1973), 997-1016.
2. Dolk, Dan: "Tobit Program Version 1.0," FORTRAN documentation, University of Arizona, (1979).
3. Fair, Ray: "A Note on the Computation of the Tobit Estimator," Econometrica, 45(1977), 1723-1727.
4. Poirier, Dale: "The Use of the Box-Cox Transformation in Limited Dependent Variable Models," Journal of the American Statistical Association, 73(1978), 284-287.
5. Tobin, James: "Estimation of Relationships for Limited Dependent Variables," Econometrica, 26(1958), 24-36.

SAS LOG OF TOBIT ANALYSIS

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1      OPTIONS NODATE NONUMBER;
2      DATA ALL NONZEROS ZEROS;
3      INPUT YEAR UM E ENC SL PR;
4      URATIO = UM / E;
5      CRATIO = SL / ENC;
6      IF (CRATIO > 0) THEN OUTPUT ALL NONZEROS;
7      ELSE OUTPUT ALL ZEROS;
8      CARDS;

```

NOTE: DATA SET WORK.ALL HAS 50 OBSERVATIONS AND 8 VARIABLES. 191 OBS/TRK.  
NOTE: DATA SET WORK.NONZEROS HAS 38 OBSERVATIONS AND 8 VARIABLES. 191 OBS/TRK.  
NOTE: DATA SET WORK.ZEROS HAS 12 OBSERVATIONS AND 8 VARIABLES. 191 OBS/TRK.  
NOTE: THE DATA STATEMENT USED 0.74 SECONDS AND 116K.

```

59      PROC MATRIX;
60      SETUP;
61      FETCH Y DATA=NONZEROS(KEEP = CRATIO);
62      FETCH NONZEROS
63      DATA=NONZEROS(KEEP = PR URATIO);
64      K = NCOL(NONZEROS) + 1;
65      R = NROW(NONZEROS);
66      X = J(R,1) || NONZEROS;
67      FETCH ZEROS
68      DATA=ZEROS(KEEP = PR URATIO);
69      S = NROW(ZEROS);
70      XBAR = J(S,1) || ZEROS;
71      FREE ZEROS NONZEROS;
72      T = R + S;
73      STEP1:
74      XCOV = INV((X') * X);
75      LSBETA = XCOV * (X') * Y;
76
77      XCOVXBAR = XCOV * (XBAR');
78      INIT:
79      IF (S > R) THEN BETA = J(K,1,0);
80      ELSE BETA = LSBETA;
81      NDRMCDEF = 1 / (SQRT(2 * ARCOS(-1)));
82      STEP2:
83      DO ITRCT = 1 TO 100;
84      LAMBDA = .50 - (INT(ITRCT / 20) * .05);
85      SIGMA2 = ((Y') * (Y - (X * BETA))) / R;
86      IF (SIGMA2 < 0) THEN SIGMA2 = .5;
87
88      SIGMA = SQRT(SIGMA2);
89      STEP3:
90      ZBAR = (XBAR * BETA) / SIGMA;
91      PHI = NORMCOEF * EXP(-.5 * (ZBAR ** 2));
92      PHICAP = PROBNORM(ZBAR);
93      GAMMABAR = PHI / (1 - PHICAP);
94      STEP4:
95      BETASQGL = LSBETA -
96      SIGMA * (XCOVXBAR * GAMMABAR);
97      BETANEW = BETA + (LAMBDA * (BETASQGL - BETA));
98      TEST:
99      IF (MAX(ABS(BETANEW - BETA)) < .001) THEN DO;
100     BETAHAT = BETANEW;
101     GO TO PREDCALC;
102     END;
103     ELSE BETA = BETANEW;
104     END;
105     OUTLOOP:
106     NOTE FAILURE TO REACH CONVERGENCE;
107     STOP;
108     PREDCALC:
109     FETCH ALL DATA=ALL(KEEP = PR URATIO)
110     COLNAME=INDNAMES;
111     XT = J(T,1) || ALL;
112     FETCH YOBS DATA=ALL(KEEP = CRATIO)
113     COLNAME=DEP VAR;
114     FETCH YID DATA=ALL(KEEP=YEAR) COLNAME=YIDNAME;
115     FREE ALL;
116     YHAT = (XT * BETAHAT);
117     S2HAT = ((Y') * (Y - (X * BETAHAT))) / R;
118     SIGMAHAT = SQRT(S2HAT);
119     SSQTOTAL = SSQ(Y) - ((SUM(Y) ** 2) / T);
120     ZT = (YHAT) / SIGMAHAT;
121     RESIDUAL = YOBS - YHAT;
122     VARGCALC:
123     PHIT = NORMCOEF * EXP(-.5 * (ZT ** 2));
124     PHICAPT = PROBNORM(ZT);
125     A = (((PHIT ** 2) / (1 - PHICAPT)) +
126     PHICAPT - (PHIT * ZT)) / S2HAT;
127     B = (((ZT ** 2) * PHIT) + PHIT -
128     ((ZT * (PHIT ** 2)) / (1 - PHICAPT)))
129     / (2 * (SIGMAHAT ** 3));
130     C = (((ZT ** 2) * (PHIT ** 2) +
131     (2 * PHICAPT)) / (1 - PHICAPT))
132     - ((ZT * PHIT) + ((ZT ** 3) * PHIT))
133     / (4 * (S2HAT ** 2));
134     SUMAXXT = J(K,K,0);

```

\* NOTATION FOLLOWS FAIR (3);  
\* Y: NONZERO OBS OF DEP VAR;  
\* K: # OF EXPLANATORY VARS;  
\* R: # OF NONZERO OBS;  
\* X: X MATRIX FOR NONZERO OBS;  
\* S: # OF ZERO OBS;  
\* XBAR: X MATRIX FOR ZERO OBS;  
\* T: TOTAL # OF OBS;  
\* LSBETA: OLSQ BETA FOR NON-ZERO OBS;  
\* INITIAL VALUE FOR BETA;  
\* ARCOS(-1) = PI;  
\* ITRCT: ITERATION COUNTER;  
\* LAMBDA: DAMPING FACTOR;  
\* SET SIGMA2 = SMALL + VALUE IF INITIAL VALUE < 0;  
\* BETASQGL: INTERMEDIATE BETA;  
\* BETANEW: NEW BETA ESTIMATE;  
\* TEST FOR CONVERGENCE;  
\* 100 ITERATION LIMIT;  
\* CALCULATE PREDICTED VALUES;  
\* XT: X MATRIX FOR ALL OBS;  
\* YOBS: DEP VAR FOR ALL OBS;  
\* YHAT: PREDICTED Y VALUES;  
\* S2HAT: EST OF SIGMA \*\* 2;  
\* SSQTOTAL: TOTAL SUM OF SQS;  
\* VAR-COV MATRIX CALCULATION;

```

135 DO I = 1 TO T;
136   SUMAXXT = SUMAXXT +
137     (A(I,1) # ((XT(I,*)') * XT(I,*)));
138 END;
139 VCOV = INV((SUMAXXT || ((XT') * B))
140           // ((B') * XT) || (SUM(C)));
141
142 PRTOUT:
143 CONSTANT = ' ';
144 NOTE CONVERGENCED IN ITRCT ITERATIONS;
145 PRINT ITRCT COLNAME=CONSTANT ROWNAME=CONSTANT;
146 PRINT FORMAT=$CHAR8. DEP_VAR COLNAME=CONSTANT
147        ROWNAME=CONSTANT;
148 CONSTANT = 'CONSTANT';
149 RLABEL1 = CONSTANT // (INDNAMES');
150 CLABEL1 = 'ESTIMATE' || 'T-STAT' ||
151           'PR > |T|' || 'STDERROR';
152 BETASE = SQRT(VECDIAG(VCOV(1:K,1:K)));
153 TSTATS = BETAHAT #/ BETASE;
154 TPROB = (1 - PROBT(ABS(TSTATS),J(K,1,R))) # 2;
155 IND_VAR = BETAHAT || TSTATS || TPROB || BETASE;
156 PRINT IND_VAR FORMAT=16.8 ROWNAME = RLABEL1
157        COLNAME = CLABEL1;
158 CONSTANT = 'SIGMA**2';
159 RLABEL2 = RLABEL1 // CONSTANT;
160 CLABEL2 = (RLABEL1') || CONSTANT;
161 NOTE VAR-COV MATRIX FOR (BETA, SIGMA**2); ;
162 PRINT VCOV FORMAT=16.8 ROWNAME=RLABEL2
163        COLNAME=CLABEL2;
164 CONSTANT = ' ';
165 PRINT SIGMAHAT S2HAT SSQTOTAL ROWNAME=CONSTANT
166        COLNAME=CONSTANT;
167 YLABEL = 'Y OBS' || 'Y HAT' || 'RESIDUAL';
168 CLABEL3 = YIDNAME || DEP_VAR || YLABEL(*,2:3);
169 YPRT = YID || YOBS || YHAT || RESIDUAL;
170 PRINT YPRT FORMAT=12.6 COLNAME=CLABEL3;
171 CLABOUT = INDNAMES || CLABEL3;
172 YOUT = XT(*,2:K) || YPRT;
173 OUTPUT YOUT OUT=TOBITOUT(DROP=ROW)
174        COLNAME=CLABOUT;
175 TITLE1 TOBIT ANALYSIS USING THE SAS MATRIX PROCEDURE;

```

NOTE: DATA SET WORK.TOBITOUT HAS 50 OBSERVATIONS AND 6 VARIABLES. 250 OBS/TRK.  
NOTE: THE PROCEDURE MATRIX USED 2.98 SECONDS AND 166K AND PRINTED PAGES 1 TO 2.

FIGURE 1. SAMPLE OUTPUT FROM TOBIT ANALYSIS PROCEDURE USING SAS

TOBIT ANALYSIS USING THE SAS MATRIX PROCEDURE

CONVERGENCED IN ITRCT ITERATIONS

ITRCT

7

DEP\_VAR

CRATIO

IND_VAR	ESTIMATE	T-STAT	PR >  T	STDERRDR
CDNSTANT	0.26765447	1.11505199	0.27183047	0.24003766
PR	-0.75376441	-0.77340500	0.44406740	0.97460504
URATIO	0.57245086	0.96618464	0.34006082	0.59248599

VAR-COV MATRIX FOR (BETA, SIGMA\*\*2):

VCOV	CONSTANT	PR	URATIO	SIGMA**2
CONSTANT	0.05761808	-0.19208061	-0.12381475	-0.00001945
PR	-0.19208061	0.94985498	0.28502116	-0.00008774
URATIO	-0.12381475	0.28502116	0.35103964	0.00004712
SIGMA**2	-0.00001945	-0.00008774	0.00004712	0.00009031

SIGMAHAT

0.330691

S2HAT

0.109357

SSQTOTAL

3.62327

YPRT	YEAR	CRATIO	Y_HAT	RESIDUAL
ROW1	21.000000	0.000000	0.293363	-0.293363
ROW2	22.000000	0.000000	0.267600	-0.267600
ROW3	23.000000	0.000000	0.411741	-0.411741
ROW4	24.000000	0.000000	0.186800	-0.186800
ROW5	25.000000	0.000000	0.205410	-0.205410
ROW6	26.000000	0.000000	0.281675	-0.281675
ROW7	27.000000	0.000000	0.196689	-0.196689
ROW8	28.000000	0.000000	0.382339	-0.382339
ROW9	29.000000	0.000000	0.245974	-0.245974
ROW10	30.000000	0.000000	0.251800	-0.251800
ROW11	31.000000	0.000000	0.304945	-0.304945
ROW12	32.000000	0.000000	0.296836	-0.296836
ROW13	33.000000	0.409091	0.432315	-0.023224
ROW14	34.000000	0.142857	0.292298	-0.149441
ROW15	35.000000	0.153409	0.204767	-0.051368
ROW16	36.000000	0.273890	0.385495	-0.111605
ROW17	37.000000	0.148571	0.309154	-0.160583
ROW18	38.000000	0.903061	0.351632	0.551430
ROW19	39.000000	0.707317	0.361164	0.346153
ROW20	40.000000	0.584906	0.271191	0.313714
ROW21	41.000000	0.573770	0.394035	0.179736
ROW22	42.000000	0.444444	0.192783	0.251662
ROW23	43.000000	0.495268	0.389293	0.105975
ROW24	44.000000	0.658654	0.313161	0.345493
ROW25	45.000000	0.529412	0.273321	0.256091
ROW26	46.000000	0.512346	0.378707	0.133639
ROW27	47.000000	0.741935	0.362298	0.379637
ROW28	48.000000	0.683495	0.441137	0.242358
ROW29	49.000000	0.694245	0.322553	0.371692
ROW30	50.000000	0.351562	0.197664	0.153899
ROW31	51.000000	0.660000	0.423777	0.236223
ROW32	52.000000	0.491228	0.300840	0.190388
ROW33	53.000000	0.746063	0.368278	0.377785
ROW34	54.000000	0.329114	0.266448	0.062668
ROW35	55.000000	0.450604	0.437701	0.012903
ROW36	56.000000	0.465257	0.223117	0.242140
ROW37	57.000000	0.316667	0.180909	0.135458
ROW38	58.000000	0.170803	0.415397	-0.244534
ROW39	59.000000	0.488764	0.249585	-0.239179
ROW40	60.000000	0.145161	0.361554	-0.216393
ROW41	61.000000	0.387782	0.418735	-0.030953
ROW42	62.000000	0.741935	0.357225	0.384711
ROW43	63.000000	0.770492	0.101254	0.669238
ROW44	64.000000	0.680851	0.249482	0.431369
ROW45	65.000000	0.145067	0.338177	-0.192109
ROW46	66.000000	0.405405	0.278563	0.126842
ROW47	67.000000	0.434959	0.412414	0.022545
ROW48	68.000000	0.427184	0.456400	-0.029215
ROW49	69.000000	0.379205	0.342545	0.036660
ROW50	70.000000	0.186047	0.242908	-0.056861